

File S1: Supporting Information

Covariance properties of the fitness distribution

To analyze the covariances of different fitness classes, we begin with Eq. (7) of the main text, which expresses $\delta x_k(\tau)$ in terms of eigenvectors $\delta x_k(\tau) = \sum_j \psi_k^{(j)} a_j(\tau)$. Projecting on the left eigenvectors then results in equations for $a_j(\tau)$

$$\frac{d}{d\tau} a_j(\tau) = -j a_j(\tau) + \sum_k \phi_k^{(j)} \sqrt{\frac{\bar{x}_k}{N_S}} \eta_k(\tau) \quad (1)$$

where $\eta_k(\tau)$ are uncorrelated Gaussian white noise terms with $\langle \eta_k(\tau) \eta_l(\tau') \rangle = \delta_{kl} \delta(\tau - \tau')$. Since each noise term η_k contributes to all a_j , the noise induces correlated fluctuations of the $a_j(\tau)$, which we need to understand in order to analyze the fluctuations of the fitness distributions. The inhomogeneous Eq. (1) has the solution

$$a_j(\tau) = \int_{-\infty}^{\tau} d\tau' e^{-j(\tau-\tau')} \sum_k \phi_k^{(j)} \sqrt{\frac{\bar{x}_k}{N_S}} \eta_k(\tau') \quad (2)$$

The autocorrelation function of the loadings of different eigendirections separated by $\Delta\tau$ in time is therefore given by

$$\begin{aligned} \langle a_i(\tau) a_j(\tau + \Delta\tau) \rangle &= \int_{-\infty}^{\tau} d\tau' \int_{-\infty}^{\tau+\Delta\tau} d\tau'' e^{-i(\tau-\tau')-j(\tau+\Delta\tau-\tau'')} \sum_{k,l} \frac{\phi_k^{(i)} \phi_l^{(j)} \sqrt{\bar{x}_k \bar{x}_l}}{N_S} \langle \eta_k(\tau') \eta_l(\tau'') \rangle \\ &= \int_{-\infty}^{\tau} d\tau' e^{-i(\tau-\tau')-j(\tau+\Delta\tau-\tau')} \sum_k \frac{\phi_k^{(i)} \phi_k^{(j)} \bar{x}_k}{N_S} \\ &= \frac{e^{-j\Delta\tau}}{i+j} \sum_k \frac{\phi_k^{(i)} \phi_k^{(j)} \bar{x}_k}{N_S} \end{aligned} \quad (3)$$

where we have used $\langle \eta_k(\tau) \eta_l(\tau') \rangle = \delta_{kl} \delta(\tau - \tau')$.

Correlation functions n_0 and the mean

To calculate the variances and covariance of x_0 and the mean fitness, we express them in terms of the eigenmodes $a_j(\tau)$ ($j > 0$)

$$\delta x_0(\tau) = \sum_{j>0} \psi_0^{(j)} a_j(\tau) = e^{-\lambda} \sum_{j>0} a_j(\tau) \quad (4)$$

$$\delta \bar{k}(\tau) = \sum_{j>0, k} k \psi_k^{(j)} a_j(\tau) = \sum_{j>0, k} k (\bar{x}_{k-j} - \bar{x}_k) a_j(\tau) = \sum_{j>0} j a_j(\tau) \quad (5)$$

The auto-correlation of x_0

The auto-correlation of x_0 is given by

$$\begin{aligned} \langle x_0(\tau) x_0(\tau + \Delta\tau) \rangle &= e^{-2\lambda} \sum_{i, j > 0} \frac{e^{-j\Delta\tau}}{i+j} \sum_k \frac{\phi_k^{(i)} \phi_k^{(j)} \bar{x}_k}{N_S} \\ &= \frac{e^{-\lambda}}{N_S} \sum_{i, j > 0} \frac{\lambda^{i+j} e^{-j\Delta\tau}}{i+j} \sum_{k=0}^{\min(i, j)} \frac{(-1)^{i+j} \lambda^{-k}}{(j-k)!(i-k)!k!} \\ &= \frac{e^{-\lambda}}{N_S} \int_0^\infty dz \sum_{i, j > 0} e^{-z(i+j)} \lambda^{i+j} e^{-j\Delta\tau} \sum_{k=0}^{\min(i, j)} \frac{(-1)^{i+j} \lambda^{-k}}{(j-k)!(i-k)!k!} \end{aligned} \quad (6)$$

Let us focus on the triple sum inside the integral and simplify it by introducing $a = -\lambda e^{-z}$ and $b = -\lambda e^{-z-\Delta\tau}$. Furthermore, let us look at the $i = j$ and the $i \neq j$ contributions separately. The diagonal contribution ($i = j$) is

$$\begin{aligned} \sum_{i=0} a^i b^i \sum_{k=0}^i \frac{\lambda^{-k}}{(i-k)!(i-k)!k!} &= \sum_{i>0} \sum_{k=0}^i \frac{(ab)^{i-k} (ab)^k \lambda^{-k}}{k!((i-k)!)^2} \\ &= \sum_{k>0} \sum_{i \geq k} \frac{(ab)^{i-k} (ab)^k \lambda^{-k}}{k!((i-k)!)^2} + \sum_{i>0} \frac{(ab)^i}{(i!)^2} \\ &= \sum_{k>0} \frac{(ab)^k \lambda^{-k}}{k!} \sum_{n \geq 0} \frac{(ab)^n}{(n!)^2} + J_0(-\iota 2\sqrt{ab}) - 1 \quad \text{using } n = i - k \\ &= (e^{ab/\lambda} - 1) J_0(-\iota 2\sqrt{ab}) + J_0(-\iota 2\sqrt{ab}) - 1 \\ &= e^{ab/\lambda} J_0(-\iota 2\sqrt{ab}) - 1 \end{aligned} \quad (7)$$

where $J_n(z)$ is the n th Bessel function of first kind, and $\iota = \sqrt{-1}$. When evaluating the off-diagonal contribution, we will encounter terms like

$$\sum_{k>0} \frac{(ab)^k}{k!(k+m)!} = \sum_k \frac{(ab)^k}{k!(k+m)!} - \frac{1}{m!} = \frac{J_m(2\iota\sqrt{ab})}{(\iota\sqrt{ab})^m} - \frac{1}{m!} \quad (8)$$

The off-diagonal contribution can be further split into the parts $i > j$ and $i < j$ which can be evaluated as follows:

$$\begin{aligned}
& \sum_{0 < i < j} a^i b^j \sum_{k \geq 0} \frac{\lambda^{-k}}{k!(j-k)!(i-k)!} = \sum_{i > 0} \sum_{j > i} a^i b^{j-i+i} \sum_{k \geq 0} \frac{\lambda^{-k}}{k!(i+(j-i)-k)!(i-k)!} \\
&= \sum_{i > 0} \sum_{m > 0} a^i b^{m+i} \sum_{k \geq 0} \frac{\lambda^{-k}}{k!(i+m-k)!(i-k)!} \quad \text{using } m = j - i \\
&= \sum_{k > 0} \sum_{m > 0} \sum_{i \geq k} \frac{a^i b^{m+i} \lambda^{-k}}{k!(i+m-k)!(i-k)!} + \sum_{m > 0} \sum_{i > 0} \frac{a^i b^{m+i}}{(i+m)!i!} \\
&= \sum_{k > 0} \sum_{m > 0} \sum_{n \geq 0} \frac{a^{n+k} b^{m+n+k} \lambda^{-k}}{k!(n+m)!n!} + \sum_{m > 0} b^m \sum_{i > 0} \frac{(ab)^i}{(i+m)!i!} \quad \text{using } n = i - k \\
&= \sum_{k > 0} \sum_{m > 0} \frac{a^k b^{m+k} \lambda^{-k}}{k!} \frac{J_m(2\sqrt{ab})}{(\sqrt{ab})^m} + \sum_{m > 0} b^m \left(\frac{J_m(2\sqrt{ab})}{(\sqrt{ab})^m} - \frac{1}{m!} \right) \\
&= \sum_{m > 0} \sum_{k \geq 0} \frac{a^k b^k b^{m/2} a^{-m/2} \lambda^{-k}}{k!} \frac{J_m(2\sqrt{ab})}{(\sqrt{ab})^m} - e^b + 1 = \sum_{m > 0} \left(\frac{b}{a} \right)^{m/2} e^{ab/\lambda} \frac{J_m(2\sqrt{ab})}{(-1)^{m/2}} - e^b + 1 \\
&= e^{ab/\lambda} \sum_{m > 0} \left(\frac{b}{a} \right)^{m/2} (-1)^m J_m(2\sqrt{ab}) - e^b + 1
\end{aligned} \tag{9}$$

The off-diagonal terms for $i > j$ is obtained by interchanging a and b such that the full off-diagonal contribution is

$$e^{ab/\lambda} \sum_{m > 0} \left[\left(\frac{b}{a} \right)^{m/2} + \left(\frac{a}{b} \right)^{m/2} \right] (-1)^m J_m(2\sqrt{ab}) - e^a - e^b + 2 \tag{10}$$

Next, we use the definition of the generating function of the Bessel functions ([?], 8.511)

$$e^{\frac{1}{2}(t-t^{-1})z} = J_0(z) + \sum_{m > 0} (t^m + (-t)^{-m}) J_m(z) \tag{11}$$

which turns the off-diagonal contribution into

$$e^{ab/\lambda} \left(e^{\sqrt{ab}(\sqrt{\frac{b}{a}} + \sqrt{\frac{a}{b}})} - J_0(2\sqrt{ab}) \right) - e^a - e^b + 2 \tag{12}$$

Combining the diagonal and off-diagonal contributions and substituting a and b , we find for the integrand in Eq. (??)

$$e^{ab/\lambda + \sqrt{ab}(\sqrt{\frac{b}{a}} + \sqrt{\frac{a}{b}})} - e^a - e^b + 1 = e^{\lambda e^{-2z} - \Delta\tau} - \lambda e^{-z} - \lambda e^{-z - \Delta\tau} - e^{-\lambda e^{-z}} - e^{-\lambda e^{-z} - \Delta\tau} + 1 \tag{13}$$

The auto-correlation of x_0 is therefore given by

$$\begin{aligned}\langle x_0(\tau)x_0(\tau + \Delta\tau) \rangle &= \frac{e^{-\lambda}}{N_s} \int_0^\infty dz \left(e^{\lambda e^{-2z-\Delta\tau}} e^{-\lambda(e^{-z}+e^{-z-\Delta\tau})} - e^{-\lambda e^{-z}} - e^{-\lambda e^{-z-\Delta\tau}} + 1 \right) \\ &= \frac{e^{-\lambda}}{N_s} \int_0^1 \frac{d\theta}{\theta} \left(e^{\lambda\theta^2 e^{-\Delta\tau} - \lambda\theta(1+e^{-\Delta\tau})} - e^{-\lambda\theta} - e^{-\lambda\theta e^{-\Delta\tau}} + 1 \right)\end{aligned}\quad (14)$$

Auto-correlation of the mean fitness

The autocorrelation function of the mean is defined as

$$\begin{aligned}\langle \delta\bar{k}(\tau)\delta\bar{k}(\tau + \Delta\tau) \rangle &= \sum_{i,j>0} ij \langle a_i(\tau)a_j(\tau + \Delta\tau) \rangle \\ &= \partial_\mu \partial_\nu \frac{1}{N_s} \int_0^\infty dz \sum_{i,j>0} \mu^i \nu^j e^{-z(i+j)} \lambda^{i+j} e^{-j\Delta\tau} \sum_{k=0}^{\min(i,j)} \frac{(-1)^{i+j} \lambda^{-k}}{(j-k)!(i-k)!k!}\end{aligned}\quad (15)$$

where the last line is to be evaluated at $\nu = \mu = 1$. Hence the problem is reduced to the one already solved with $a = -\mu\lambda e^{-z}$ and $b = -\nu\lambda e^{-z-\Delta\tau}$. We find

$$\langle \delta\bar{k}(\tau)\delta\bar{k}(\tau + \Delta\tau) \rangle = \frac{\lambda e^\lambda}{N_s} \int_0^1 d\theta e^{-\Delta\tau} e^{\lambda\theta^2 e^{-\Delta\tau} - \lambda(1+e^{-\Delta\tau})\theta} (\theta + \lambda\theta (\theta e^{-\Delta\tau} - 1) (\theta - 1)) \quad (16)$$

Cross-correlation of x_0 and the mean fitness

When calculating the cross-correlation between x_0 and the mean fitness we have to distinguish the cases where x_0 precedes the mean fitness and vice-versa. Otherwise, the calculation proceeds almost unchanged from the cases discussed above.

$$\begin{aligned}\langle x_0(\tau)\delta\bar{k}(\tau + \Delta\tau) \rangle &= \sum_{i,j>0} j \langle a_i(\tau)a_j(\tau + \Delta\tau) \rangle \\ &= \partial_\nu \frac{1}{N_s} \int_0^\infty dz \sum_{i,j>0} a^i b^j \sum_{k=0}^{\min(i,j)} \frac{(-1)^{i+j} \lambda^{-k}}{(j-k)!(i-k)!k!}\end{aligned}\quad (17)$$

with $a = -\lambda e^{-z}$, $b = -\nu\lambda e^{-z-\Delta\tau}$ if $\Delta\tau > 0$ and $a = -\lambda e^{-z+\Delta\tau}$, $b = -\nu\lambda e^{-z}$ if $\Delta\tau < 0$. The result is

$$\langle \delta x_0(\tau)\delta\bar{k}(\tau + \Delta\tau) \rangle = \frac{\lambda}{N_s} \begin{cases} e^{-\Delta\tau} \int_0^1 d\theta \left((\theta - 1)e^{e^{-\Delta\tau}\lambda\theta^2 - \lambda(1+e^{-\Delta\tau})\theta} + e^{-e^{-\Delta\tau}\lambda\theta} \right) & \Delta\tau > 0 \\ \int_0^1 d\theta \left((e^{\Delta\tau}\theta - 1)e^{e^{\Delta\tau}\lambda\theta^2 - \lambda(1+e^{\Delta\tau})\theta} + e^{-\lambda\theta} \right) & \Delta\tau < 0 \end{cases}\quad (18)$$

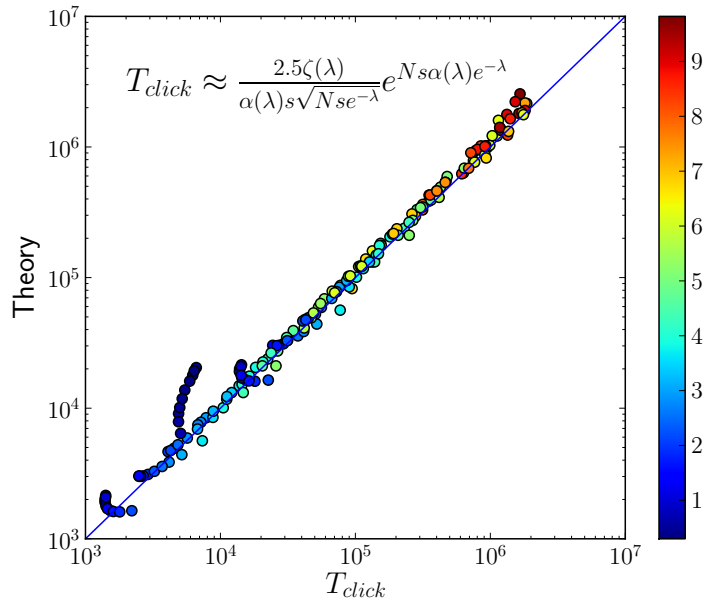


Figure S 1: The approximation of the mean time between clicks of the ratchet is accurate over a large range of parameters if $Ns\alpha(\lambda)e^{-\lambda}$ is large compared to one. $Ns\alpha(\lambda)e^{-\lambda}$ determines whether the clicks of the ratchet are far apart compared to the relaxation time of the distribution and is indicated as the color of the data points. The condition $Ns\alpha(\lambda)e^{-\lambda} > 1$ is violated for the fastest clicks shown, resulting in the deviation of the dark blue points.